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## Long-term thermal sensitivity of Earth's tropical forests

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## 1 Abstract

2 The sensitivity of tropical forest carbon to climate is a key uncertainty in predicting global climate 3 change. While short-term drying and warming are known to impact forests it is unknown if such effects 4 translate into long-term responses. Here we analyse 590 permanent plots measured across the tropics to 5 derive the equilibrium climate controls on forest carbon. Maximum temperature is the most important predictor of aboveground biomass (-9.1 Mg C ha<sup>-1</sup> °C<sup>-1</sup>), primarily by reducing woody productivity, and 6 7 with a greater rate of decline in the hottest forests (>32.2 °C). Our results nevertheless reveal greater 8 thermal resilience than observations of short-term variation imply. To realise the long-term climate 9 adaptation potential of tropical forests requires both protecting them and stabilising the Earth's climate.

10

One sentence summary. Biome-wide variation in tropical forest carbon stocks and dynamics shows
long-term thermal resilience.

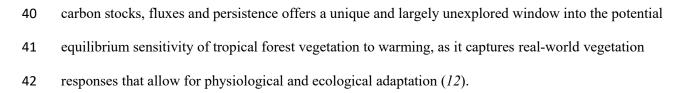
### 13 <u>Main text</u>

The response of tropical terrestrial carbon to environmental change is a critical component of global 14 climate models (1). Land-atmosphere feedbacks depend on the balance of positive biomass growth 15 stimulation by  $CO_2$  fertilisation (i.e.  $\beta$ ) and negative responses to warmer temperatures and any 16 change in precipitation (i.e.  $\gamma$ ). Yet the climate response is so poorly constrained that it remains one of 17 the largest uncertainties in Earth system models (2, 3), with the temperature sensitivity of tropical land 18 carbon stocks alone differing by  $> 100 \text{ Pg C} \circ \text{C}^{-1}$  among models (2). Such uncertainty impedes our 19 understanding of the global carbon cycle, limiting our ability to simulate the future of the Earth 20 21 system under different long-term climate mitigation strategies. A critical long-term control on tropical 22 land-atmosphere feedbacks is the sensitivity to climate ( $\gamma$ ) of tropical forests, where c. 40 % of the world's vegetation carbon resides (4). 23

The sensitivity to environmental change of tropical biomass carbon stocks, their rate of production and their persistence, can all be estimated by relating their short-term and inter-annual responses to variation in climate (5-7). These sensitivities are then used to constrain longer-term projections of climate responses (2). Such approaches typically find that higher minimum temperatures are strongly associated with slower tree growth and reduced forest carbon stocks, likely due to increased respiration at higher temperatures (7-9). Tropical forest carbon is also sensitive to precipitation (10), with, for example, elevated tree mortality occurring during drought events (11).

31 Yet the sensitivity of ecosystems to inter-annual fluctuations may be an unreliable guide to their longer-term responses to climate change. Such responses will also be influenced by physiological 32 acclimation (12), changes in demographic rates (13), and shifts in species composition (14). For 33 34 example, both respiration and photosynthesis can acclimate under sustained temperature increases 35 (15-17), and tropical trees exhibit physiological plasticity (18) and shifts in species composition (14)36 under sustained drought. These processes could mean that tropical forests are less sensitive to climate than estimates derived from inter-annual variability imply. An alternative, complimentary approach to 37 assessing sensitivity to climate is to measure and analyse spatial variation in tropical ecosystems 38 across climate gradients as a space-for-time substitution. Such biome-wide spatial variation in forest 39

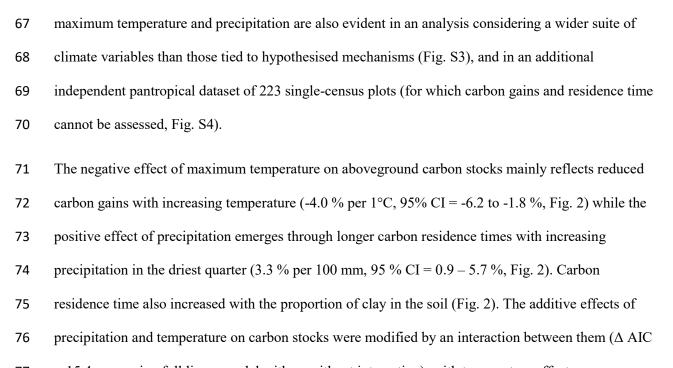
#### Sullivan et al. Thermal sensitivity of tropical forests



43 To assess the long-term climate controls on tropical forest growth and carbon stocks, here we have assembled, measured, and analysed a pan-tropical network of 590 permanent, long-term inventory 44 plots (Fig. 1, see Figs. S1-2 for ability to capture biome climate space). Our analysis combines 45 46 standardised measurements from across South American, African, Asian and Australian tropical 47 lowland forests (273, 239, 61 and 17 plots respectively). For every plot we calculated aboveground 48 carbon stocks (19). Then, to better assess the dynamic controls on aboveground carbon stocks, we also computed the rate of carbon gained by the system (aboveground woody carbon production, 49 calculated as tree growth plus newly recruited trees, in Mg C ha<sup>-1</sup> yr<sup>-1</sup>), and the carbon residence time 50 in living biomass (calculated as the ratio of living C stocks to C gains, in years). 51

52 We find considerable variation in biomass carbon among continents, with lower stocks per unit area 53 in South America compared with the Paleotropics even after accounting for environmental variables (Fig. 1). Continents with high carbon stocks had either large carbon gains (Asia), or long carbon 54 55 residence times (Africa, Fig. 1). Because of these differences among continents, which are potentially 56 due to differences in evolutionary history (20), we analyse the environmental drivers of spatial 57 variation in carbon stocks while accounting for biogeographical differences. We fitted linear models with explanatory variables representing hypothesised mechanistic controls of climate on tropical 58 forest carbon (Table S1). We also included soil covariates, continent intercepts and eigenvectors 59 describing spatial relationships amongst plots to account for other sources of variation (21). 60

Forest carbon stocks were most strongly related to maximum temperature (-5.9 % per 1°C increase in maximum temperature, 95 % CI = -8.6 to -3.1 %, Fig. 2, equivalent to -9.1 Mg C ha<sup>-1</sup> °C<sup>-1</sup> for a stand with the mean carbon stocks in our dataset, 154.6 Mg C ha<sup>-1</sup>), followed by rainfall (+2.4 % per 100 mm increase in precipitation in the driest quarter, 95 % CI = 0.6 - 4.3 %, Fig. 2, equivalent to 0.04 Mg C ha<sup>-1</sup> mm<sup>-1</sup> for a stand with the mean carbon stocks in our dataset), with no statistically significant relationship with minimum temperature, wind speed or cloud cover (Fig 2). The effects of



77 = 15.4 comparing full linear model with or without interaction), with temperature effects more

negative when precipitation is low (Fig. S6). The interaction was through shortening carbon residence time ( $\Delta AIC = 11.9$ ) rather than reducing carbon gains (model without interaction better,  $\Delta AIC =$ 

80 1.4).

An alternative analysis using decision tree algorithms (22) also showed maximum temperature and precipitation to be important (Fig. S7). This decision tree approach, which can capture complex nonlinear relationships (22), indicated potential non-linearity in the relationships between carbon stocks and both temperature and precipitation, with the positive effect of increasing dry season precipitation on residence times strengthening when precipitation was low, and the negative effect of maximum temperature intensifying at high temperatures (Fig. S7).

We further investigated non-linearity in the temperature relationship using breakpoint regression (supported over linear regression based on lower AIC,  $\Delta$  AIC = 15.0), which revealed that above 32.2 % CI = 31.7 – 32.6 °C) the relationship between carbon stocks and maximum temperature

90 became more negative (cooler than breakpoint: -3.8 % °C<sup>-1</sup>, warmer than breakpoint: -14.7 % °C<sup>-1</sup>,

91 Fig. 3). By partitioning carbon stocks into their production and persistence we find that this non-

92 linearity reflects changes to carbon residence time ( $\Delta AIC = 10.6$ ) rather than gains ( $\Delta AIC = 1.7$ ).

93 Overall, our results thus indicate two separate climate controls on carbon stocks: a negative linear

effect of maximum temperature through reduced carbon gains, and a non-linear negative effect of
maximum temperature, ameliorated by high dry-season precipitation, through reduced carbon
residence time.

97 The effect of temperature on carbon residence time only emerges when dry season precipitation is low so is consistent with theoretical expectations that negative effects of temperature on tree longevity are 98 exacerbated by moisture limitation, rather than being independent of it and due to increased 99 100 respiration costs alone (23). This could occur through high vapour pressure deficits in hot and dry 101 forests increasing mortality risk by causing hydraulic stress (23, 24), or carbon starvation due to 102 limited photosynthesis as a result of stomatal closure (23). Notably, the temperature-precipitation interaction we find for aboveground stocks is in the opposite direction to temperature-precipitation 103 interactions reported for soil carbon. In soils, moisture limitation suppresses the temperature response 104 105 of heterotrophic respiration (25), while in trees moisture limitation enhances the mortality risks of 106 high temperatures.

107 The temperature effects on biomass carbon stocks and gains are primarily due to maximum rather 108 than minimum temperature. This is consistent with high daytime temperatures reducing  $CO_2$ assimilation rates, for example due to increased photorespiration or longer duration of stomatal 109 closure (26, 27), whereas if negative temperature effects were to have increased respiration rates there 110 111 should be a stronger relationship with minimum (i.e. night-time) temperature. Critically, minimum temperature is unrelated to above ground carbon stocks both pan-tropically and in the one continent, 112 South America, where maximum and minimum temperature are largely decoupled (r = 0.33; Fig. S8). 113 While carbon gains are negatively related to minimum temperature (Fig S9) this bivariate relationship 114 115 is weaker than with maximum temperature, and disappears once the effects of other variables are accounted for (Fig. 2). Finally, in Asia, the tropical region which experiences the warmest minimum 116 temperatures of all, both carbon stocks and carbon gains are highest (Fig. 1, Fig. S11). 117

Overall our results suggest that tropical forests have considerable potential to acclimate and adapt to the effects of night-time minimum temperatures, but are clearly sensitive to the effects of daytime maximum temperature. This is consistent with ecophysiological observations suggesting that the

121 acclimation potential of respiration (15) is greater than that of photosynthesis (17). The temperature sensitivity revealed by our analysis is also considerably weaker than short-term sensitivities 122 associated with inter-annual climate variation (8). For example, by relating short-term annual climate 123 anomalies to responses in plots, the effect of a 1°C increase in temperature on carbon gains has been 124 125 estimated as more than three-fold our long-term, pantropical result (28). This stronger long-term thermal resilience is likely due to a combination of individual acclimation and plasticity (15-17), 126 differences in species' climate responses (29) leading to shifts in community composition due to 127 128 changing demographic rates (12) and the immigration of species with higher performance at high 129 temperatures (12).

Our pantropical analysis of the sensitivity to climate of aboveground forest carbon stocks, gains and 130 persistence shows that warming reduces carbon stocks and gains from woody productivity. Using a 131 reference carbon stock map (30) and applying our estimated temperature sensitivity (including non-132 133 linearity) while holding other variables constant leads to an eventual biome-wide reduction of 14.1 Pg C in live biomass (including scaling to estimate carbon in roots) for a 1°C increase in maximum 134 temperature (95 % CI = 6.9 - 20.7 Pg). This compares with a large range of projected sensitivities in 135 the subset of coupled climate carbon cycle models that report vegetation carbon  $(1 - 58 \text{ Pg C} \circ \text{C}^{-1})$ , 136 137 although we note that these models have not been run to equilibrium (see SI Methods).

138 Our results suggest that stabilising global surface temperatures at 2°C above pre-industrial levels will cause a potential long-term biome-wide loss of 35.3 Pg C (95 % CI = 20.9 - 49.0 Pg, estimates with 139 alternative baseline biomass maps 24.0 - 28.4 Pg, Fig. S12). The greatest long-term reductions in 140 carbon stocks are projected in South America, where baseline temperatures and future warming are 141 142 both highest (Fig. 4, Fig. S13). This warming would push 71 % of the biome beyond the thermal threshold - maximum temperature of 32.2°C - where larger long-term reductions in biomass are 143 expected (Fig. S14). Of course, growth stimulation by carbon dioxide (31) will partially or wholly 144 offset the effect of this temperature increase, depending on both the level of atmospheric carbon 145 dioxide that limits warming to 2°C above pre-industrial levels and the fertilization effect of this 146 carbon dioxide on tropical trees. Although CO<sub>2</sub> fertilisation is expected to reduce temperature induced 147

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carbon losses from biomass across the tropics (Table S3), our analysis indicates that CO<sub>2</sub> fertilisation 148 is not enough to offset long-term temperature induced carbon losses within Amazonia (Fig. S15). 149 The long-term climate sensitivities derived from our pan-tropical field measurements incorporate 150 ecophysiological and ecological adaptation, and so provide an estimate of the long-term, quasi-151 152 equilibrium, response of tropical vegetation to climate. We note that this thermal adaptation potential may not be fully realised in future responses because (i) the speed of temperature rises may exceed 153 species' adaptive capabilities, (ii) habitat fragmentation may limit species' ability to track changes in 154 the environment, and (iii) other human impacts such as logging and fire can increase the vulnerability 155 156 of forest carbon stocks to high temperatures. While many tropical forests are under severe threat of 157 conversion, our results show that, in the long-run, tropical forests that remain intact can continue to store high levels of carbon under high temperatures. Achieving the biome-wide climate resilience 158 potential we document depends on limiting heating and on large-scale conservation and restoration to 159 160 protect biodiversity and allow species to move.

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# 412 Supplementary Materials:

- 413 Materials and Methods
- 414 Figures S1-S15
- 415 Tables S1-S3
- 416 References (32-85)

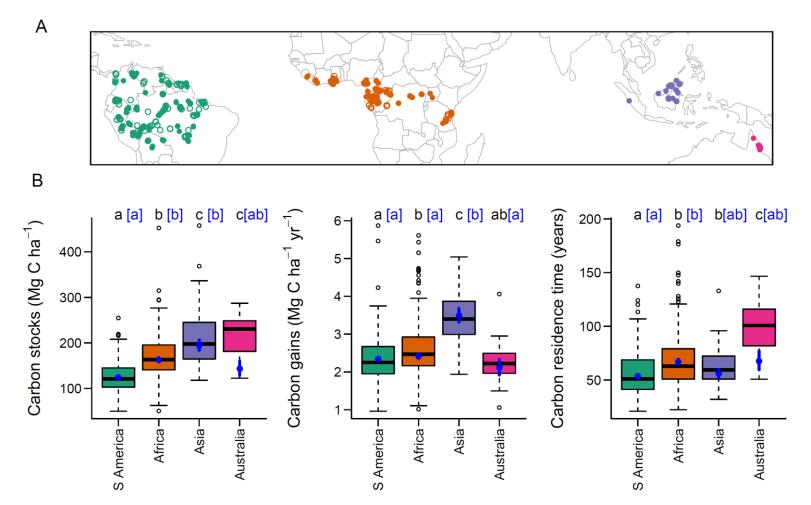
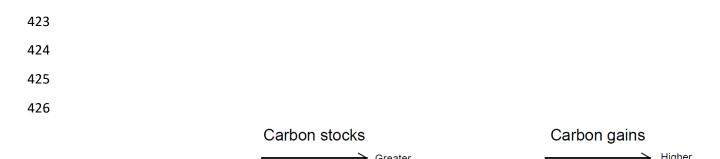
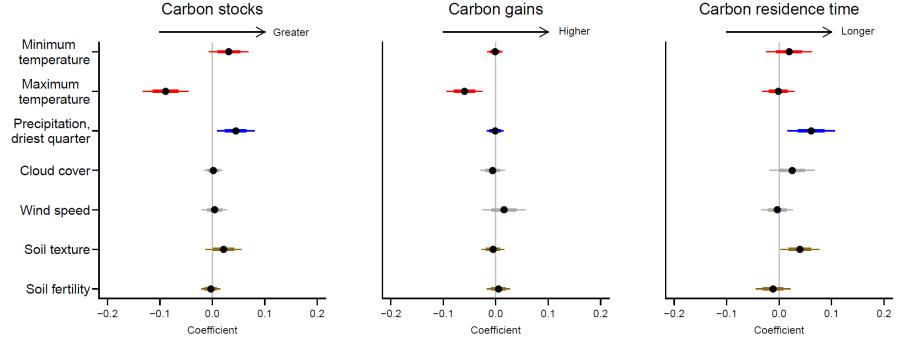


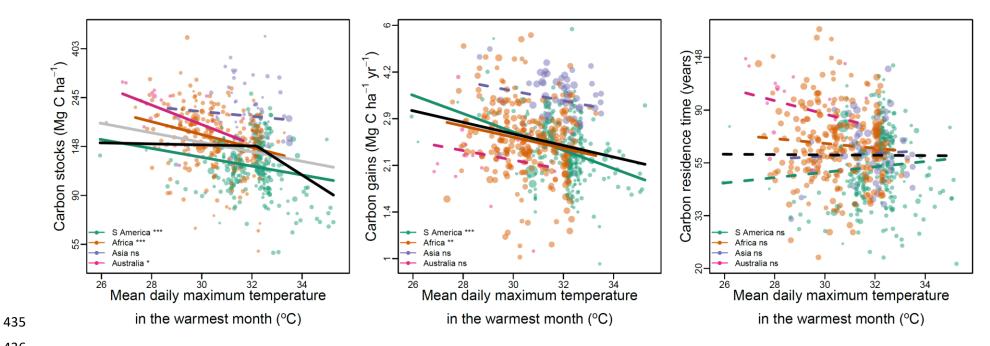
Figure 1. Spatial variation in tropical forest carbon. (A) Our plot network. Filled symbols show multi-census plots used in the main analysis, open symbols show single-census plots used as an independent dataset. (B) Variation in carbon among continents. Boxplots show raw variation while blue points show estimated mean values ( $\pm$  SE) after accounting for environmental variation. Letters denote statistically significant differences between continents (P < 0.05) based on raw data (black) or accounting for environmental effects (blue, square brackets).





427

Figure 2. Correlates of spatial variation in tropical forest carbon. Points show coefficients from model-averaged general linear models. Variables that did not occur in well-supported models are shrinkage adjusted towards zero. Coefficients are standardised so that they represent change in the response variable for one standard deviation change in the explanatory variable. Error bars show standard errors (thick lines) and 95% confidence intervals (thin lines). Soil texture is represented by the percentage clay, and soil fertility by cation exchange capacity. The full models explained 44.1 %, 31.4 % and 30.9 % of spatial variation in carbon stocks, gains and residence time respectively. Coefficients are shown in Table S2. Results are robust to using an alternative allometry to estimate tree biomass (Fig. S5).



437 Figure 3. Temperature effects on tropical forest carbon stocks, carbon gains from woody productivity and carbon residence time. Black lines show the best pan-tropical relationships accounting for environmental covariates. The grey line shows the additional linear pan-tropical relationship for carbon stocks. 438 Coloured lines show bivariate relationships within each continent. Statistically significant relationships are shown with solid lines, non-significant with 439 dashed lines. Note that the y-axis is on a log-scale. Symbol point size is proportional to weights used in model fitting based on plot size and monitoring 440 length, see SI Materials and Methods. For stocks and gains linear and break-point pan-tropical relationships are all statistically significant (P < 0.001), as are 441 442 better sampled continents. For carbon residence time, relationships with temperature are non-significant but there is a statistically significant interaction between maximum temperature and precipitation in the driest quarter (Figure S6). Relationships with other variables are shown in Fig. S8-S10. \*\*\* P < 443 0.001, \*\* P < 0.01, \* P < 0.05, ns  $P \ge 0.05$ 444



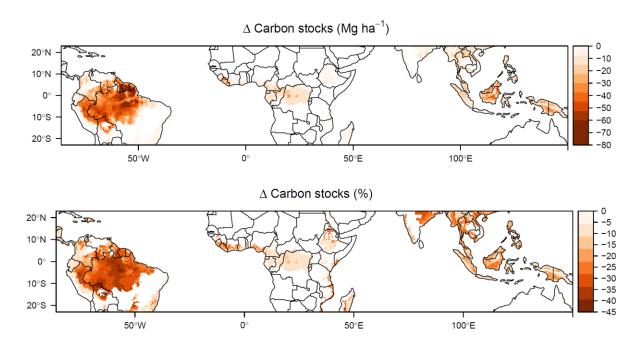


Figure 4. Long-term change in carbon stocks due to global surface temperature warming of
approximately 2°C. Maps show the predicted absolute and relative change in tropical forest carbon
stocks. Note that parts of the biome become warmer than currently observed in our dataset (Fig. S14).
See Fig. S12 for predictions using alternative carbon reference maps. Predictions are based on
temperature alone and do not include precipitation changes (for which future patterns of change are
uncertain) or potential moderation via elevated CO<sub>2</sub> (see Fig. S15 for analysis incorporating this).

## 455 Supporting information for Long-term thermal Sensitivity of the Earth's Tropical

#### 456

#### Forests

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Supporting information for Sullivan et al.

# 505 This file includes:

- 506 Materials and Methods
- 507 Figures S1 S15
- 508 Tables S1 S2

Supporting information for Sullivan et al.

#### 510 Materials and Methods

## 511 Forest census data

512 Our plots come from the RAINFOR, AfriTRON, and T-FORCES networks. Forest inventory plots

513 were located in lowland (<1200 m), old-growth, closed-canopy forests that were not known to have

been subject to anthropogenic disturbance through fire or selective logging. Plots characterised

- floristically as dry forest were not included, as were plots that received less than 1200 mm
- 516 precipitation each year. We also did not include plots in white sand, swamp and seasonally flooded
- 517 forests, as we expect these to experience marked edaphic constraints (extreme nutrient limitation for
- 518 white sand forests (32), stress caused by hypoxic conditions for swamp and seasonally flooded forests
- 519 (33)). All plots were  $\ge 0.2$  ha (median size = 1 ha) and were monitored for at least two years (median
- 520 monitoring period = 9.7 years). All censuses were prior to the 2015-16 very strong El Niño event, as
- 521 we expected that event to supress carbon gains relative to the long-term mean.
- 522 Forest inventory plots were sampled using standardised protocols (34), where all live stems with

523 diameter  $\geq$  100 mm were measured at 1.3 m or 50 cm above buttresses and deformities. Trees were

524 tagged so that the same tree could be identified in subsequent censuses. In some cases the point of

- 525 diameter measurement (POM) had to be moved due to upward growth of buttresses and deformities.
- 526 For these trees we use the  $D_{mean}$  approach from Talbot et al. (35).

527 In a few cases (6 plots) the minimum diameter measured changed over time, or palms and

528 *Phenakospermum* were excluded in some censuses. For these, we estimated aboveground biomass

- 529 (AGB, subsequently converted to carbon stocks) and aboveground woody production (AGWP,
- subsequently converted to carbon gains) using a minimum diameter or taxonomic protocol that could
- be consistently applied across censuses, and scaled these values by the aboveground biomass ratio
- between that protocol and all stems  $\geq 100$  mm protocol for censuses when all stems were measured.

533 Some plots had nested designs where the plot was split into subplots with different minimum diameter

- protocols (69 plots). For these, we only analysed the area conforming to our minimum diameter
- protocol. For analysis, we grouped small ( $\leq 0.5$  ha) plots within 1 km of each other, and also grouped
- 536 contiguous larger plots (18 plots), as these will experience equivalent climate and larger plots are less
- 537 sensitive to stochastic tree fall events (*36*).
- 538 Data were curated in ForestPlots.net (37, 38), or were subject to equivalent offline handling, and
- 539 experienced the same quality control procedures. Details of quality control procedures are described
- 540 in Brienen et al. (39). Our final dataset consists of 590 sampling units (hereafter plots) covering 637.2
- ha, with 2.2 million measurements of 670,499 unique stems. For validating models of carbon stocks
- an additional dataset of 223 single-census plots using the same measurement protocols was assembled
- 543 from the same networks (see section "Validation with independent single-census plot dataset" below).



## 545 <u>Estimating above-ground biomass</u>

546 Diameter measurements were converted to estimates of aboveground biomass (AGB). For dicot trees547 we used the allometric equation

[1]

[3]

548 AGB = 
$$0.673 \times (\rho D^2 H)^{0.976}$$
,

from Chave et al. (40), where  $\rho$  is wood density (from (41, 42)) and H is tree height estimated using

allometric equations described below. For monocots and tree ferns, we used a palm-specific

allometric equation

552  $\ln(AGB) = -3.3488 + 2.7483.\ln(D),$  [2]

from Goodman et al. (43), where D is the measured diameter.

The heights of a subset of trees in our dataset were measured in the field, either with a laser 554 rangefinder, hypsometer, or clinometer, or directly by climbing the tree. We filtered this dataset to 555 stems with measured diameters, height < 90 m, diameters > 90 mm DBH, as height-diameter 556 allometries of saplings differ from those of more mature trees, and to stems that were not broken, 557 558 leaning or fallen. This gave a total of 78,899 height measurements. We used this dataset to fit local height-diameter allometric models, as these refine AGB estimates by capturing spatial variation in 559 height-diameter allometries missed by large-scale allometric models (44). Height data were not 560 561 available from every plot, so to ensure consistent treatment of plots height-diameter models were 562 constructed for each biogeographic region. We fitted three parameter asymptotic models (45) of the 563 form

564 
$$H = a(1 - \exp(-bD^{C})),$$

where a, b and c are estimated parameters ('Weibull' models, 46). We fitted these models either 565 treating each observation equally or with case weights proportional to each trees' basal area. These 566 weights give more importance to large trees during model fitting. We selected the best fitting of these 567 568 models, determining this as the model that minimised prediction error of stand biomass when 569 calculated with estimated heights or observed heights (44). Weibull models were implemented using the nls function in R with default settings. Starting values of a = 25, b = 0.05 and c = 0.7 were chosen 570 following trial and error as they led to regular model convergence. Where models did not converge 571 572 this was usually because the height-diameter relationship did not reach an asymptote, so in these cases 573 we used the log-log model  $\ln(H) = a + b(\ln(D))$  to estimate height, where b gives the scaling exponent 574 of a power law relationship between height and diameter. We checked if models gave unrealistic 575 predictions by applying models to predict the height of all trees in the biogeographic region, and

excluded models that predicted any tree height 10 % higher than the tallest tree we recorded in thatcontinent.

578

### 579 Estimating above-ground woody production

580 We estimated AGWP following Talbot et al. (35). AGWP is comprised of four components, (1) the

sum of growth of surviving trees, (2) the sum of AGB of new recruits, (3) the sum of unobserved

growth of trees that died during a census interval and (4) the sum of growth of unobserved recruits

that entered then died during a census interval. Accounting for the latter two components is necessary
to avoid census-interval length effects, as more AGWP in these components will be missed due to the

585 greater mortality of trees that accumulates over longer census intervals.

586 Components 3 and 4 can be estimated using two quantities that can be calculated from observed stem-

dynamics in each plot; per-area annual recruitment  $(R_a)$  and per-capita annual mortality  $(m_a)$ . Per-

588 capita mortality is calculated from the ratio of surviving stems to initial stems, using equation 5 in

589 Kohyama et al. (47). Per-area annual recruitment is calculated using estimated mortality rates and the

590 observed change in the number of stems over a census interval, using equation 11 of Kohyama et al.

591 (47).

To estimate the unobserved growth of stems that died during a census interval, we first use plot-level per-capita mortality rates  $(m_a)$  to estimate how many trees are expected to have died in each year of the census interval, and from that calculate the mean number of years that trees that died during the census interval would have lived before death. The diameter of tree at death  $(D_{death})$  can then be estimated as

597 
$$D_{\text{death}} = D_{\text{start}} \times G \times Y_{mean}$$
 [4]

where  $D_{\text{start}}$  is the diameter at the start of the census interval, *G* is the plot-level median growth rate of the size class the tree was in at the start of the census interval (size classes are defined as D < 200 mm,  $400 \text{ mm} > D \ge 200 \text{ mm}$ , and  $D \ge 400 \text{ mm}$ ) and  $Y_{mean}$  is the mean number of years trees survived in the census interval before dying. The diameter at death is then converted to AGB at death using allometric equations (equation 1, except for ferns and monocots where equation 2 is used), and the unobserved growth is calculated as the difference between AGB at death and AGB at the start of the census.

- To estimate the growth of recruits that were not observed because they died during the census
- 606 interval, we first need to estimate the number of unobserved recruits. This can be estimated from per-
- area annual recruitment ( $R_a$ ) and per-capita annual mortality ( $m_a$ ):  $R_a$  gives the number of stems per ha
- that recruit in a given year, and the probability of each recruit surviving until the next census  $(P_{surv})$  is

- 609  $P_{surv} = (1-m_a)^T$ , where T is the number of years remaining in the census interval. The number of
- 610 recruits in a given year that survive to the next census is  $R_a P_{surv}R_a$ . Summing this for each year in a
- 611 census interval gives the total number of unobserved recruits in that census interval. We then need to
- estimate how long each recruit was alive for. From  $m_a$  we can calculate the number of recruits in a
- 613 given year that died in each subsequent year, and from this calculate the mean life-span of recruits in a
- 614 given year that died before the next census. The average life-span of unobserved recruits ( $Y_{mean-rec}$ ) is
- the weighted mean of each cohort's lifespan, weighted by the number of unobserved recruits in each
- 616 year. Diameter at death is given in mm by

617 
$$D_{\text{death}} = 100 + (G \times Y_{mean-rec})$$
 [5]

- 618 where G is the plot-level median growth rate of the smallest size class (i.e.  $D \le 200$  mm).
- 619 Aboveground biomass of recruits at the time of death is estimated using equation 1. These corrections
- 620 for unobserved growth have a marginal impact on AGWP calculations, collectively accounting on
- 621 average for just 2.3 % of estimated plot-level AGWP.
- 622 AGB was calculated for each census, and AGWP was calculated for each census interval, and the
- time-weighted mean of each was taken to give one value per plot. We used a time-weighted mean to
- 624 give greater importance to AGB estimates separated by longer census-intervals, as these will be more
- 625 independent. Estimates of AGB and AGWP were converted to carbon stocks and carbon gains by
- 626 multiplying by 0.456 (48). Carbon residence time was then estimated as carbon stocks /carbon gains,
- and represents the length of time carbon resides in living biomass before being passed to the litter and
- 628 necromass pools (49). Calculations to estimate AGB and AGWP were performed using the R package
- 629 BiomasaFP (*50*).
- 630

#### 631 Obtaining environmental data

Most climate data were obtained from climate data from Worldclim2 (51) as it provides the highest 632 resolution (~1 km) pantropical climate data, although we note that some regions, such as central 633 Africa, have limited station data. We extracted monthly data for the following variables: mean daily 634 minimum temperature, mean daily maximum temperature, precipitation, solar radiation and wind 635 speed, In addition to calculating the standard series of 19 bioclimatic variables, using the dismo R 636 package (52), we calculated 1) mean daily maximum temperature, BIO1 + BIO2/2, 2) mean daily 637 minimum temperature, BIO1 – BIO2/2, 3) maximum cumulative water deficit as the minimum across 638 the year of monthly cumulative water deficit W, 639

640  $W_i = W_{i-1} - \min(0, P_i - 100),$  [6]

- 641 where *P* is monthly precipitation in mm, and 100 represents measured evapotranspiration. This
- 642 calculation was run for a year from the wettest month in the year, starting at a water deficit of zero, 4)
- 643 the number of months where monthly cumulative water deficit was negative, 5) the number of months
- 644 where monthly precipitation was below 100 mm (i.e. less than evapotranspiration), 6) mean annual
- solar radiation, 7) mean annual wind speed, and 8) vapour pressure deficit (VPD = SVP vapour
- 646 pressure, where saturated vapour pressure, SVP, =  $0.611 \times e^{(17.502 \text{ temperature})/(\text{temperature} + 240.97)}$ ). We also
- 647 obtained data on cloud frequency at  $\sim$ 1 km resolution from Wilson & Jetz (53), who processed twice-
- 648 daily MODIS satellite images. Temperature values were adjusted for differences in altitude between
- the plot and the 1 km grid cell used for Worldclim interpolation, as these can differ in topographically
- diverse regions, using lapse rates, so that  $T_{plot} = T_{worldclim} + 0.005 \times (A_{worldclim} A_{plot})$ , where T is
- temperature (°C) and A is altitude (m). Temperature values were also corrected for systematic
- warming trends. To do this, the mean annual temperature in each grid-cell in each year was extracted
- from the CRU TS 3.24 dataset (54), and robust linear regression used to estimate grid-cell specific
- 654 warming rates. These were used to adjust Worldclim2 temperature values for the difference between
- the midpoint of plot monitoring and the midpoint of the Worldclim2 climatology.
- Data on soil texture and chemistry was obtained at 1 km resolution from the SoilGrids dataset (55),
- with this resolution selected to match the resolution of the climate data. From this we extracted CEC,
- 658 representing soil fertility, and percentage clay, representing soil texture. For each soil variable we
- 659 calculated the depth-weighted average for 0 30 cm.
- 660 <u>Statistical analysis</u>
- We used linear models to relate carbon, carbon gains and carbon residence time to environmental
  explanatory variables. The role of different explanatory variables was assessed using multi-model
  inference.
- Response variables were positively skewed and had positive mean-variance relationships, so were log-transformed to meet the assumption of normality and reduce heterogeneity in variances. The lognormal nature of forest carbon stocks and dynamics means that there is greater potential for variation when forests are large, which could be due to the non-linear scaling of tree biomass and tree basal area.
- We selected explanatory variables to represent hypothesised ways in which climate could affectcarbon stocks (Table S1). We assessed colinearity within this set of explanatory variables using
- 671 variance inflation factors (VIF) and pairwise correlations. Because of colinearity, we had to exclude
- 672 VPD, total precipitation, use only one of MCWD and precipitation in the driest quarter, and could
- 673 include both minimum and maximum temperature but not mean annual temperature. We used
- 674 precipitation in the driest quarter rather than MCWD as the latter is zero truncated and so is less
- amenable to regression analysis. After removing these variables all pairwise correlations (including

with soil explanatory variables) were weak enough not to cause problems through collinearity (r < 0.6and VIF < 3).

To account for variation other than in climate we also included soil variables relating to texture (% clay) and fertility (CEC), and included continent specific intercepts to account for biogeographic variation in carbon. To account for unmeasured environmental gradients (e.g. soil variation not captured by the SoilGrids variables), we used Moran's eigenvector maps as explanatory variables, selecting eigenvectors the corresponded to positive spatial autocorrelation in the distance matrix (*56*). These variables act as a proxy for unmeasured spatial gradients by capturing positive spatial associations between plots.

Plots differed in their area and the length of time they were monitored for. This is likely to affect the 685 variance of carbon stocks, carbon gains and carbon residence time, as smaller plots or plots only 686 monitored for short periods are more likely to be sensitive to the mortality of a few large trees. To 687 account for this, we used case weights relating to plot area and monitoring period. Following Lewis et 688 689 al. (57), we selected weights by relating residuals from our linear models to plot area and to plot 690 monitoring period, and subsequently assessing which root transformation of plot area/ monitoring 691 period removed the pattern in the residuals when used as a weight. Selected weights were: carbon stocks, Area <sup>1/3</sup>; carbon gains, Monitoring length <sup>1/7</sup>; carbon residence time, Area <sup>1/9</sup> + Monitoring 692 length <sup>1/12</sup> -1. 693

694 We fitted all subsets of the general linear model with explanatory variables described above, forcing 695 spatial eigenvectors into all models. We then averaged the subset of models where  $\Delta$  AIC < 4, using 696 full averaging so variables that do not appear in the model get the value of zero for their coefficients. 697 This means that model averaged coefficients of terms with limited support exhibit shrinkage towards 698 zero. Multi-model inference was performed using the MuMIn R package (*58*).

We assessed whether the two climate variables found to have important additive effects on carbon stocks in this analysis (mean daily maximum temperature in the warmest month and precipitation in the driest quarter) interacted with each other by adding an interaction term between these variables to the full generalised linear model of carbon stocks as a function of other climate and soil variables, continent and spatial eigenvectors. We compared these two models using AIC. We repeated this with carbon gains and carbon residence time as response variables.

To assess whether the temperature carbon relationship was non-linear we used breakpoint regression implemented in the segmented R package (*59*). This estimates a breakpoint in the explanatory variable at which the slope of the relationship with the response variable changes. We estimated the breakpoint for the mean daily maximum temperature in the warmest month variable in the full model with a temperature-precipitation interaction described above. We assessed the support for the breakpoint by Supporting information for Sullivan et al.

- comparing the AIC of the model with a breakpoint with the AIC of a model with a linear relationship.
- 711 We repeated this with carbon gains and carbon residence time as response variables.
- 712 We also analysed spatial variation in carbon stocks as a function of the above climate and soil
- variables and spatial eigenvectors using Random Forest decision tree algorithms (22) implemented
- via using the randomForest R package (60). We assessed variable importance by calculating the average
- increase in node purity across all decision trees (measured by residual sum of squares) when using the
- variable to split the data. We assessed modelled relationships between response and explanatory
- variables using partial plots, which show predicted change in the response variable, averaged across
- trees, when changing the explanatory variable and holding all other variables constant.
- 719 To compliment this analysis based on relationships expected *a priori*, we also performed an
- exploratory analysis to assess whether other climate variables excluded from the full general linear
- models had an effect on carbon. To do this, we fitted linear models to assess the bivariate relationship
- of carbon with each climate variable, with continent also included as an explanatory variable to
- 723 account for biogeographic variation in forest characteristics.
- 724

## 725 Validation with independent single-census plot dataset

726 We assessed whether the relationships with environmental variables identified in the analyses of

- 727 multi-census plot data described above held when applied to an additional dataset of 223 single-
- census plots. As the single-census data were not used in any of the analyses above they did not
- influence modelling decisions, so provide an independent test of the relationships identified with the
- 730 multi-census plot analysis.
- 731 Single-census plots were extracted from the ForestPlots.net database (*37, 38*) using the same plot-
- selection criteria as for the multi-census plots, except that censuses during or following the 2015-16
- strong El Niño were included in the single-census plot dataset as we expected that carbon stocks,
- value value
- 735 We fitted a general linear model with the five climate explanatory variables, soil fertility and texture,
- continent and spatial eigenvector, and model averaging of all subsets of this model as described for
- the multi-census plots. We performed this analysis using just the single-census plots and a combined
- 738 dataset of single and multi-census plots.
- 739
- 740
- 741

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## 742 <u>Scaling results to the biome</u>

743 We applied the non-linear relationship between carbon stocks and mean daily maximum temperature in the warmest month identified by the breakpoint regression to estimate the total change in carbon 744 stock due to temperature effects alone for different scenarios of temperature increase. We delimited 745 the biome extent using the WWF tropical and subtropical moist broadleaved forest biome (61), 746 restricted to tropical latitudes, and further refined it by excluding grid-cells with < 50 Mg C ha<sup>-1</sup> using 747 data from (30), as these are unlikely to be forest. Calculations were conducted at 10-minute 748 749 resolution. The non-linear relationship between temperature and carbon means that the change in biomass for a given increase in temperature will depend on the baseline temperature. For each grid-750 751 cell we predicted the percentage change in carbon for a given temperature increase from the baseline 752 temperature in that grid-cell based on the non-linear relationship identified in our statistical model, 753 holding all other variables constant. We then used a reference carbon stock map (30) to convert 754 percentage change to change in carbon stocks per hectare (in Mg ha<sup>-1</sup>). To calculate change in carbon 755 stocks for the whole grid-cell, we multiplied change per hectare by the area of the grid-cell in 756 hectares, and then adjusted this by the proportion of the grid-cell that was forested by multiplying by 757 2014 forest cover (62). Total change for the biome (in Pg) was calculated by summing these grid-cell 758 level values. Uncertainty due to our statistical model was assessed by generating multiple predictions 759 by resampling model parameters (breakpoint threshold, slope below breakpoint, slope above 760 breakpoint), and extracting quantiles from the resultant distribution of predicted change values. Aboveground biomass carbon values were scaled to include root biomass based on a root to shoot 761

- ratio of 0.19 in tropical evergreen forests (63).
- 763 The Avitabile et al (*30*) aboveground biomass map was chosen to provide reference carbon stocks.
- 764 While other maps have previously been produced by Saatchi et al. (64) and Baccini et al. (65) we
- selected the Avitabile map because it synthesises the earlier maps (see Mitchard et al. (66) for
- 766 discussion of substantial differences between these maps) and is anchored by more field data.
- 767 Importantly, the Avitabile map reproduces spatial patterns in aboveground biomass that have been
- 768 described from field data but are absent in the Saatchi or Baccini maps, including the much higher
- biomass density of north-east Amazonian forests due to tall trees and very high wood density (67).
- 770 Nevertheless, we also investigated the consequences of using the Saatchi or Baccini maps for our
- estimates of biomewide thermal sensitivity and spatial patterns of change in carbon stocks (Fig S15).
- 772 We investigated three temperature change scenarios. Firstly, we applied a 1°C increase to all
- 1.5°C above pre-
- industrial levels for the equilibrium temperature response of tropical forest carbon. Finally, we
- assessed the consequence of global temperatures stabilizing 2°C above pre-industrial levels. For the
- 1776 latter two we obtained data from CMIP5 climate models, using downscaled future climate projections

777 based on the Worldclim climatology (68). As downscaling was performed using Worldclim version 778 1.4 (69) and our statistical models use Worldclim version 2, we calculated the warming anomaly in 779 each grid-cell from the current Worldclim version 1.4 conditions, and applied this to the Worldclim 2 780 data to obtain future temperature. RCP scenarios and time-points were chosen to give global temperature increases that best match 1.5°C and 2°C above pre-industrial. Importantly, these future 781 climate projections were used to capture the spatially varying nature of warming, and our predictions 782 relate to the long-term response of vegetation if the climate stabilised at these new warming levels, 783 rather than being predictions of transient responses at these specific time-points. For 1.5°C we used 784 RCP 2.6 averaged for 2040-2060 (median temperature increase across models =  $1.5^{\circ}$ C, (70)). For 785  $2^{\circ}$ C, we used RCP 2.6 averaged for 2040-2060 (median temperature increase models =  $1.9^{\circ}$ C (70)). 786 Note that predicted increases in maximum temperatures were often considerably greater than the 787 global increase, especially in South America. For both scenarios we used the median predicted 788 temperature change for each grid-cell from an ensemble of 15 models (BCC-CSM1-1, CCSM4, 789 CNRM-CM5, GFDL-CM3, GFDL-ESM2G, GISS-E2-R, HadGEM2-AO, HadGEM-ES, IPSL-790 791 CM5A-LR, MIROC-ESM-CHEM, MIROC-ESM, MIROC5, MPI-ESM-LR, MRI-CGCM3,

792 NorESM1-M).

We assessed the potential for long-term carbon dioxide growth stimulation to offset these long-term
 temperature effects. We used CO<sub>2</sub> concentrations from the RCP scenarios and time-points described

above, which approximate the long-term concentrations if the climate stabilised at the new

temperatures (71). Thus the 1.5°C and 2°C scenarios were associated with CO<sub>2</sub> concentrations of 443

ppm and 487 ppm respectively (72). We cannot assess the effect of CO<sub>2</sub> on biomass from our spatial

dataset, so instead used independent estimates of CO<sub>2</sub> effects from other sources. Firstly, we obtained

CO<sub>2</sub> only effects on net primary production (NPP) extracted from an ensemble of CMIP5 earth system

800 models by (73). This gives the proportional change in NPP for evergreen forests (note that this also

801 includes boreal forests) over 1980-2010, standardised to a 100 ppm increase in  $CO_2$  concentration. To

propagate this through to changes in AGB under future CO<sub>2</sub> conditions we first estimated the

803 logarithmic dependency of NPP on  $CO_2$  (74) by substituting values of NPP and  $CO_2$  at time zero and *t* 804 (from (73)) into the equation,

805 
$$NPP_t = NPP_0 \left[ 1 + \beta \ln \left( \frac{|CO_2|_t}{|CO_2|_0} \right) \right]$$
 Equation 7

- 806 This equation can be used to compute NPP annually given an initial NPP estimate and a time series of
- atmospheric  $CO_2$  concentrations (from a combination of the observed record from pre-industrial and
- the RCP 4.5 scenario, modified so that it stabilises at 487 or 443 ppm depending on warming
- scenario). Initial pre-industrial NPP was back-calculated from present-day values using Equation 7,
- 810 with 13.3 Mg C ha<sup>-1</sup> yr<sup>-1</sup> (mean of nine Amazon plots where NPP has been measured, from (75)) used

for present-day NPP. To propagate NPP into change in woody biomass (following (49)) we used theequation

813 
$$\frac{dM_{\text{wood}}}{dt} = \alpha_{\text{wood}} N_{\text{P}} - \frac{M_{\text{wood}}}{\tau_{\text{wood}}}$$
Equation 8

814 where  $M_{wood}$  is woody biomass,  $N_p$  is NPP,  $\alpha_{wood}$  is the allocation of NPP to wood (taken as 0.33, the 815 mean value across nine plots from (75)) and  $\tau_{wood}$  is the residence time of woody biomass, taken as 816 59.1 years (the median value across plots used in this study). This model (equations 7 and 8) was run 817 from pre-industrial to 2500, enabling us to see the equilibrium effect of increased CO<sub>2</sub> concentrations 818 on biomass, assuming temporally invariant allocation and residence time. We calculated the 819 proportional change in biomass from 2000 to 2500, and applied this to the reference carbon stock map 820 to obtain predicted equilibrium change in aboveground biomass due to CO<sub>2</sub> effects.

821 The effects of CO<sub>2</sub> in earth system models have been reported to be larger than those deduced from 822 satellite data or  $CO_2$  enrichment experiments (73), so we also ran the above model using changes in 823 NPP reported from a synthesis of free-air  $CO_2$  enrichment experiments conducted in forests (73). Finally, we looked at the impact of using CO<sub>2</sub> effects derived from a recent large meta-analysis of 824 CO<sub>2</sub> enrichment experiments (76), which reported a 12.5 % increase in biomass of tropical trees for a 825 826 250 ppm increase in  $CO_2$  concentration. As this relationship was reported to be linear (76) we used 827 linear interpolation to estimate the change in biomass under CO<sub>2</sub> concentrations associated with each warming scenario (i.e. 443 and 487 ppm). To estimate long-term changes in biomass accounting for 828 both temperature and carbon dioxide, we first applied the  $CO_2$  relationship to estimate the change in 829 830 biomass due to carbon dioxide growth stimulation, and then assessed the effects of warmer temperatures from this revised baseline. Our approach allows a simple assessment of CO<sub>2</sub> effects 831 exploring a range of different effect strengths. Real-world responses will likely be more complex, 832

- 833 with, for example, nutrient limitation potentially affecting the extent to which growth is stimulated by
- 834 CO<sub>2</sub> (*76*).

## 835 <u>Temperature sensitivity of CMIP5 models</u>

836 The temperature sensitivity  $(\gamma_{LT})$  of coupled climate carbon cycle models can be identified by

837 comparing responses of carbon stocks in coupled and uncoupled simulations forced with a 1%

- 838 increase in CO<sub>2</sub> concentrations per year (respectively, these are the 1pctCO2 and esmFixClim
- simulations), following Wenzel et al. (77). Both coupled and uncoupled simulations are exposed to
- 840 the same increase in  $CO_2$  concentration, but in the uncoupled simulation temperature is not directly

841 affected by this increase in CO<sub>2</sub>.

842 Vegetation carbon outputs are reported from six CMIP5 models, each with coupled and uncoupled
843 simulations (78, 79). For all simulations, we calculated the change in vegetation carbon (the *cVeg*

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- variable) in the tropics between year 110 and year 30 of the experiment, and also calculated the
- 845 difference in land temperature (the *tas* variable). The change in vegetation carbon due to temperature
- alone was calculated by taking the difference in change in vegetation carbon in the coupled ( $\Delta C veg_C$ )
- and uncoupled ( $\Delta C veg_U$ ) simulation, and this was then divided by the change in tropical land
- 848 temperature ( $\Delta T$ ) to obtain the temperature sensitivity of the model,

849  $\gamma_{LT} = (\Delta C veg_C - \Delta C veg_U) / \Delta T$ 

Equation 9.

- 850 We calculated the temperature sensitivity of the six CMIP5 models that report vegetation carbon:
- 851 CESM-1-BGC ( $\gamma_{LT}$  = -0.7 Pg C °C<sup>-1</sup>), GFDL-ESM2M ( $\gamma_{LT}$  = -58.4 Pg C °C<sup>-1</sup>), HadGEM2-ES ( $\gamma_{LT}$  = -
- 852 9.2 Pg C °C<sup>-1</sup>), IPSL-CM5A-LR ( $\gamma_{LT}$  = -11.3 Pg C °C<sup>-1</sup>), MPI-ESM-LR ( $\gamma_{LT}$  = -22.8 Pg C °C<sup>-1</sup>) and
- 853 NorESM1-ME ( $\gamma_{LT}$  = -1.0 Pg C °C<sup>-1</sup>). Note that the simulations do not run to equilibrium (77), so
- changes in carbon stocks due to increased temperature may not be fully realised.

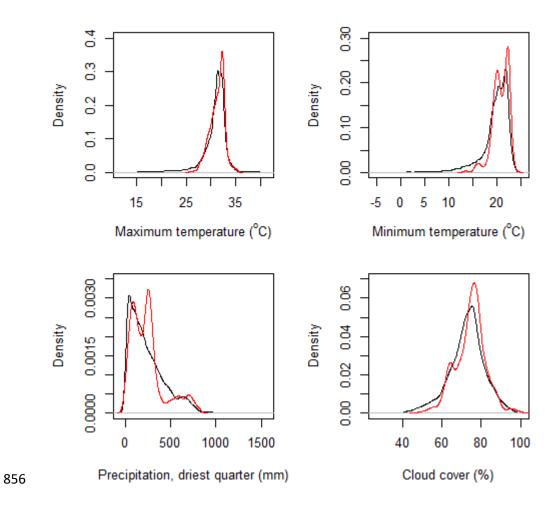


Figure S1. Climate space represented by our plot network. Red lines show the probability density
function of each variable in our multi-census plot network. Black lines show the probability density
across 10 minute grid-cells in the biome, restricted to areas with forest cover in GLC 2000 (80).



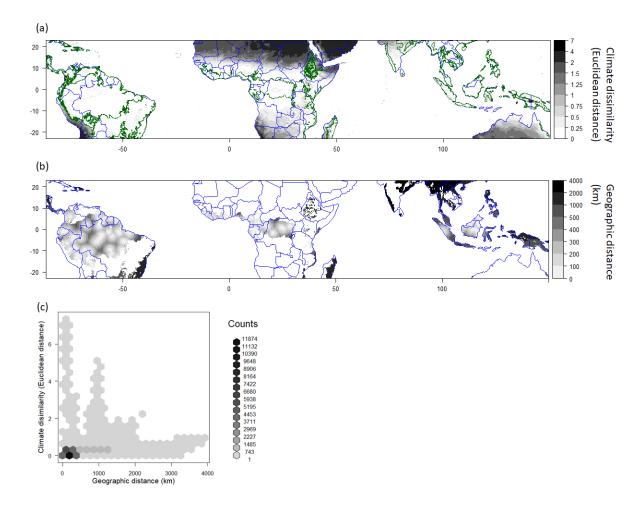


Figure S2. Ability of our plot network to represent the climate conditions found in the moist tropical 863 forest biome. (a) Minimum climate dissimilarity (measured as Euclidean distance on variables scaled 864 by their standard deviation). Climate variables used are the same as in Fig. 2) between 10 minute grid 865 866 cells and the multi-census plot network. Green lines indicate the extent of the biome. (b) Geographic distance (km) between grid cells and the multi-census plot network. (c) Relationship between climatic 867 and geographic distance of 10 minute grid cells across the tropical forest biome to our plot network. 868 869 The lack of relationship between climate dissimilarity and geographical distance, alongside the mostly low climatic dissimilarities, shows that our sampling is sufficient to capture the environmental space 870 of the biome and that we can reasonably extrapolate to geographically distant areas from our plots, 871 872 which are in any case largely deforested already and hence contribute very little to our projected biome-wide carbon response to climate change. (These tropical moist forest areas that are poorly 873 874 sampled and largely lost include the Atlantic Forests in Brazil, Andean Forests in western South 875 America, eastern Caribbean, Madagascar, and much of tropical South Asia, south China, continental 876 Southeast Asia, Philippines, Sumatra and Java).



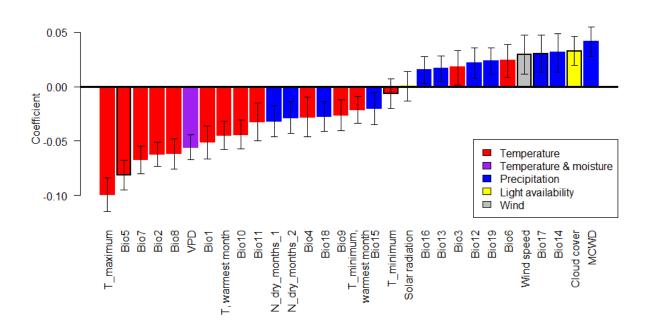
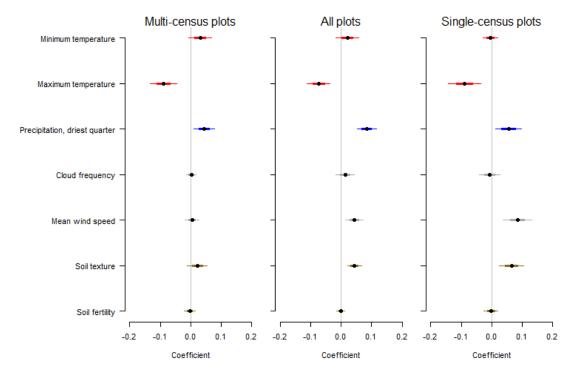


Figure S3. Relationships between individual climate variables and tropical forest aboveground carbon 879 880 stocks. Standardised coefficients are from models with the climate variable and continent as explanatory variables and show change in ln(carbon) for a standard deviation change in the 881 explanatory variable. Error bars show standard errors. Variables used in the main analysis have black 882 883 outlines. Full variable names are: T maximum – mean daily maximum temperature, Bio5 – mean daily maximum temperature in the warmest month, Bio7 – annual temperature range, Bio2 – mean 884 diurnal temperature range, Bio8 – mean temperature in the wettest quarter, VPD – vapour pressure 885 deficit, Bio1 – mean annual temperature, Bio10 – mean temperature in the warmest quarter, Bio11 – 886 mean temperature in the coldest quarter, N dry months 1 - number of months with negative 887 cumulative water deficit, N dry months 2 – number of months where precipitation is less than 888 evapotranspiration, Bio4 - temperature seasonality, Bio18 - precipitation in the warmest quarter, Bio9 889 - mean temperature in the driest quarter, T minimum warmest month - mean daily minimum 890 891 temperature in the warmest month, Bio15 – precipitation seasonality, T minimum – mean daily minimum temperature, Bio16 – precipitation in the wettest quarter, Bio13 – precipitation in the 892 wettest month, Bio3 - isothermality, Bio12 - annual precipitation, Bio19 - precipitation in the coldest 893 894 quarter, Bio6 – mean daily minimum temperature in the coldest month, Wind speed – mean daily wind speed, Bio17 - precipitation in the driest quarter, Bio14 - precipitation in the driest month, 895 Cloud cover - proportion of MODIS passes with cloud present, MCWD - maximum cumulative 896 water deficit (note this is negative when water deficit is high, so a positive relationship with MCWD 897 898 indicates higher carbon when water deficits are less).

899





903 Figure S4. Validation of tropical forest carbon stock sensitivity model against an independent dataset 904 of 223 single-census plots from our networks measured with the same protocols. Model-averaged shrinkage adjusted coefficients from multiple regression models of biomass carbon stocks as a 905 906 function of climate, soil, biogeography and spatial eigenvectors. Models were either fitted to the 907 multi-census plot dataset (as in Fig. 2), to the single-census plot dataset, or to the combined dataset. 908 This analysis shows that the relationships identified to be most important in the main multi-census 909 plot analysis (i.e. the negative relationship between carbon stocks and maximum temperature and positive relationship with precipitation in the driest quarter) are also found in an independent dataset, 910 911 which was not used for preliminary analysis so did not influence the choice of explanatory variables. 912

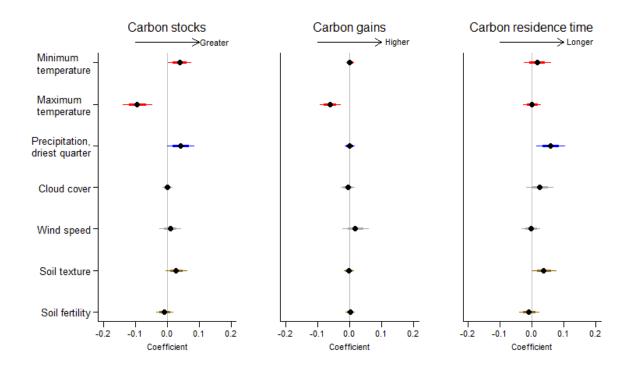
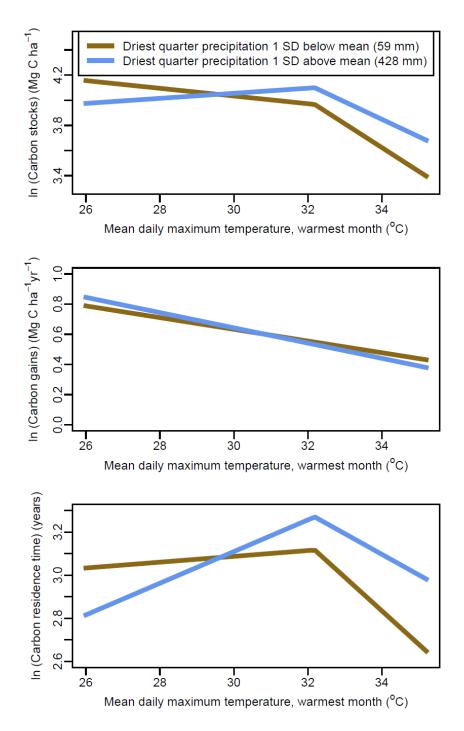




Figure S5. As Figure 2, but with aboveground biomass estimated using the Chave et al. 2005 (81)
moist forest allometric equation, which does not include a height term and is instead based on a thirdorder polynomial relationship between diameter and aboveground biomass. This indicates that our

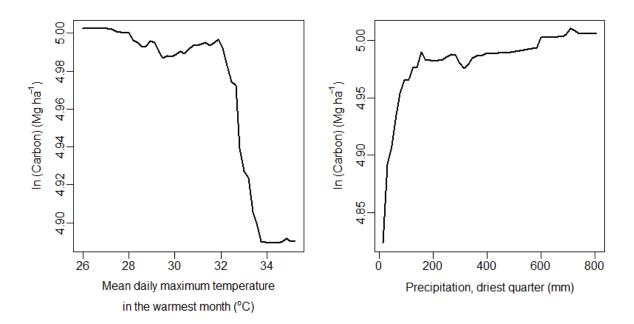
917 results are robust to using an alternative allometry to estimate aboveground biomass.





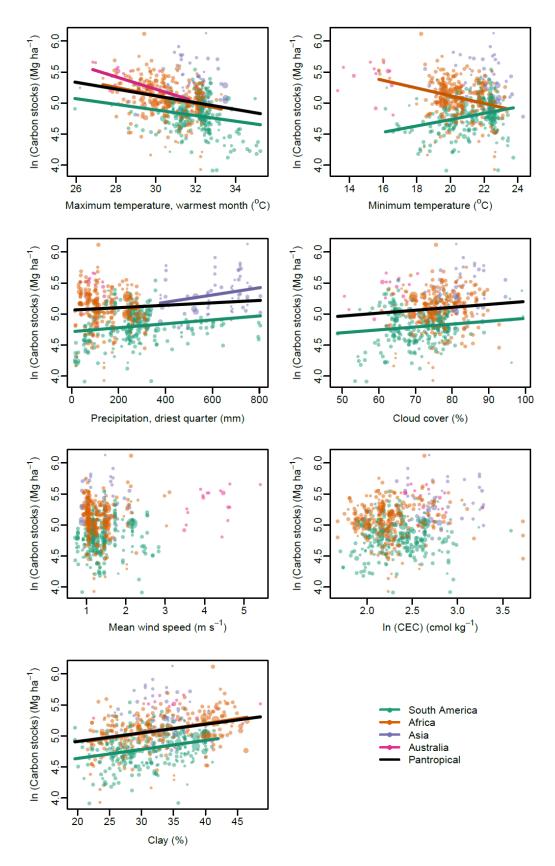


922 Figure S6. Interaction between mean daily maximum temperature in the warmest month and precipitation in the driest quarter in determining aboveground tropical forest carbon stocks, gains and 923 924 residence time. Modelled relationships with temperature are shown holding precipitation either one 925 standard deviation above or below the mean. Models with breakpoints are shown for carbon stocks 926 and residence time as they were found to be better supported based on lower AIC ( $\Delta AIC > 2$ ). Note 927 that the temperature-carbon relationship is steeper when precipitation is low for carbon stocks and 928 (above the breakpoint threshold) carbon residence time, but does not change with precipitation for 929 carbon gains. Response curves are predicted with continent set as Africa.



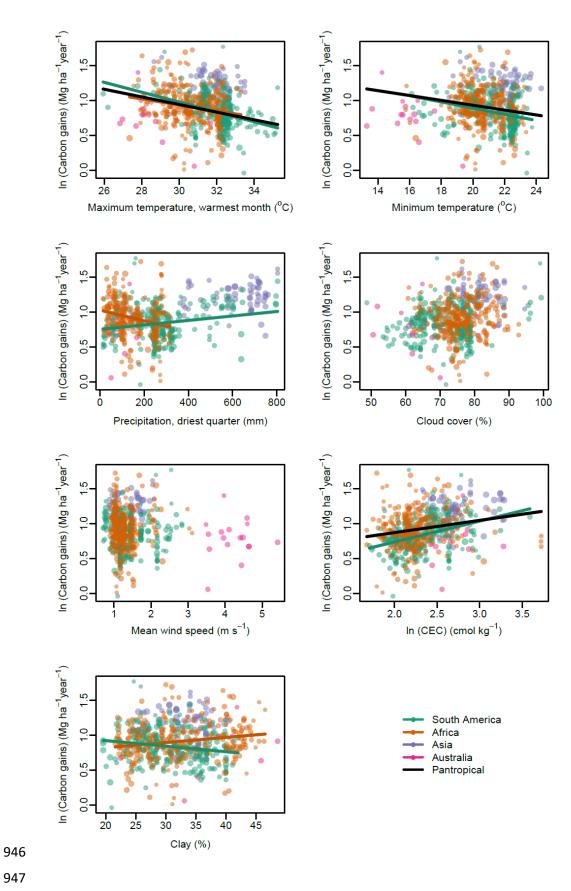


932 Figure S7. Partial relationships between tropical forest carbon stocks and the two climate variables 933 identified to be most important by the random forest decision tree algorithm. Partial plots show predicted values of carbon stocks averaged across an ensemble of decision tree models when 934 935 changing the explanatory variable of interest and holding other variables constant. The importance of 936 variables in random forest analysis is assessed by calculating the average increase in node purity 937 across all decision trees (measured by residual sum of squares) when using the variable to split the data. Higher values indicate greater importance. Maximum temperature increased node purity by 4.8 938 939 and precipitation by 4.7. For all other climate variables increases in node purity were < 3.5.

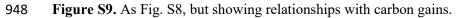


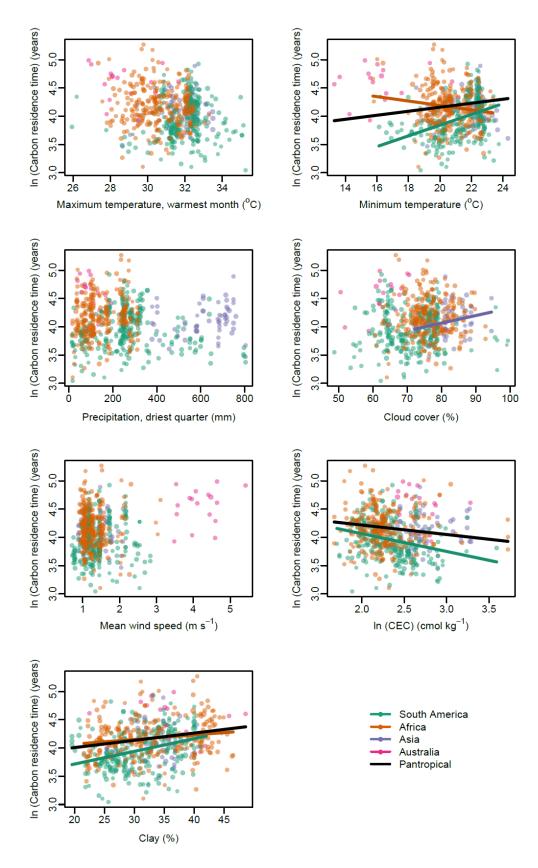


942 Figure S8. Relationships between aboveground tropical forest carbon stocks and environmental
943 predictors. Symbols and colours as in Fig. 3. Coloured lines show bivariate relationships in each
944 continent, and black lines show pan-tropical relationships also accounting for the effect of continent.
945 Lines are only plotted where statistically significant.









950 Figure S10. As Fig. S8, but showing relationships with carbon residence time.

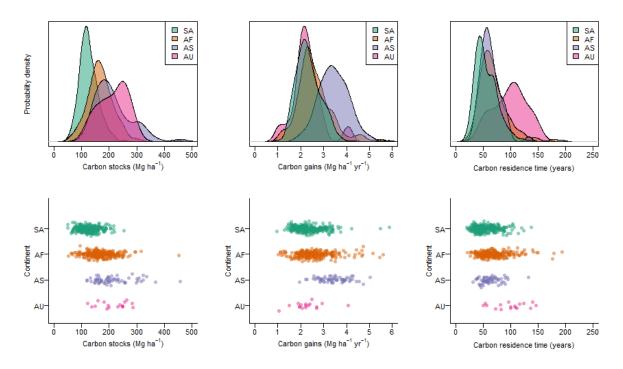
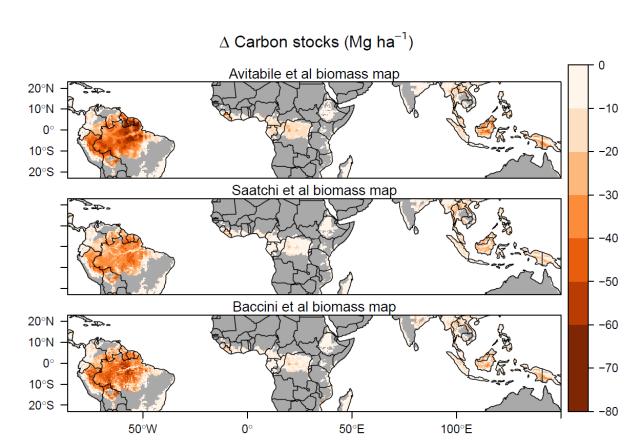




Figure S11. Variation in tropical forest aboveground carbon stocks, gains and residence time within
and amongst continents. Data are presented as empirical probability density functions (top row) and
dot-plots showing raw data points for all our multi-census plots (bottom row). SA = South America,
AF = Africa, AS = Asia, AU = Australia.

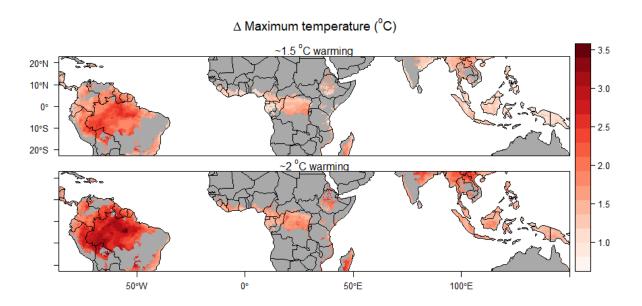


**Figure S12.** Effect of using earlier biomass reference maps for estimates of change in long-term carbon stocks for global temperature increases of ~2°C. Using aboveground biomass stock maps from Saatchi et al. (64) and Baccini et al. (65) predicted biome-wide reductions in biomass carbon stocks are 24.0 Pg (95 % CI = 5.8 - 39.6) and 28.4 Pg (95 % CI = 16.1 - 37.5) respectively. Under the ~

963  $1.5^{\circ}$ C warming scenario these are 18.4 Pg (5.8 – 30.5) and 21.1 Pg (10.2 – 29.4) respectively. Results

964 in the main text use the 2016 Avitabile et al. baseline map (30) – see methods for justification.





969	Figure S13. Biome-wide change in mean daily maximum temperature in the warmest month from
970	present conditions (based on the Worldclim climatology, 1970-2000), given global increases in
971	temperature of approximately 1.5°C and 2°C above pre-industrial levels. These levels of global
972	temperature increase are obtained from, respectively, RCP 2.6, 2040-2060 and RCP 4.5, 2040-2060 to
973	represent the potential spatial pattern of warming associated with global temperatures stabilising at
974	these levels. Global temperature increases of 1.5 and 2°C above pre-industrial levels (so ~0.8 °C and
975	~1.3 °C above our current baseline climate) would lead to mean increases in maximum temperature in
976	the warmest month across the tropical forest biome of 1.9°C and 2.4°C the current baseline climate
977	respectively.

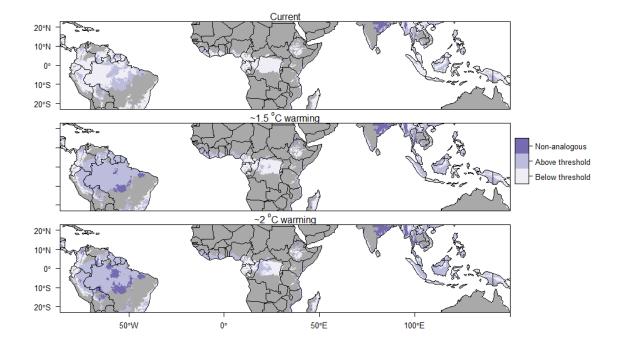
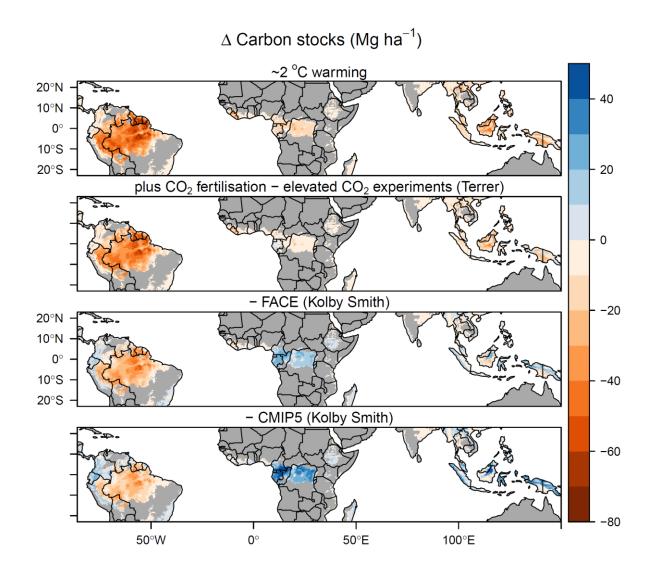


Figure S14 Areas of the biome above or below the 32.2°C threshold, above which carbon stocks
decline more rapidly with temperature, under current conditions and two warming scenarios (see Fig.
4). Areas warmer than any currently observed in our dataset (35.2°C) are also shown (non-analogous conditions). Note that even the 1.5°C warming scenario pushes most South American forests above
the 32.2°C threshold.



994 Figure S15. Predicted long-term change in aboveground carbon stocks under ~ 2°C global warming, 995 based on either temperature effects alone or when also accounting for carbon dioxide growth stimulation. CO2 fertilisation effects on equilibrium biomass levels were obtained from a recent 996 997 synthesis of results of elevated CO<sub>2</sub> experiments (Terrer et al. (76)), free-air CO<sub>2</sub> enrichment (FACE) 998 experiments (Kolby Smith et al. (73)) and CMIP5 earth system models (Kolby Smith et al. (73)). Depending on their strength, CO<sub>2</sub> effects either partially or fully ameliorate the biome-wide negative 999 1000 effects of increasing temperatures on biomass carbon stocks (Table S3), but these carbon stocks are 1001 predicted to decline over much of Amazonia even under the strongest CO2 effect considered.

1002

**Table S1.** Climate variables selected for analysis and mechanisms by which they can affect carbon stocks.

Variable selected for analysis	Mechanism to affect carbon stocks
Maximum temperature in the warmest month <sup>1</sup>	High daytime temperatures exceed photosynthesis optima (82), increase evaporative stress, causing stomatal closure and reducing time for photosynthesis (26) and increase risk of mortality through hydraulic failure and/or carbon starvation (23).
Mean daily minimum temperature	Respiration rate increases with temperature so proportion of carbon taken through photosynthesis that is allocated to wood should decline with temperature ( $83$ ). Increased respiration cost could also reduce tree longevity ( $23$ ). As respiration occurs day and night, and photosynthesis only in the day, nighttime temperature should better reflect respiration effects and daytime temperature better reflect photosynthesis effects.
Precipitation in the driest quarter <sup>2</sup>	Moisture availability could limit photosynthesis and hence carbon gains, with stomata closing when moisture availability is limiting. The risk of mortality through hydraulic failure or carbon starvation is higher when moisture is limiting (23), and this could also set a limit on potential tree size and hence tree longevity.
Cloud frequency	Increased photosynthesis and hence AGWP when light availability is greatest (i.e. cloud cover is low) $(84)$ . Alternatively, light availability could have a negative effect due to high evapotranspiration stress when cloud cover is low.
Mean wind speed	Carbon stocks are expected to be lower where physical damage through wind throw or breakage is higher, as carbon is removed more quickly from the system through mortality (85). But there is potential for greater carbon gains if forests are more dynamic.
	bio5) was selected instead of mean daily maximum temperature as it was more strongly decoupled from t some of these effects, but was too strongly correlation with maximum temperature to include as an
1	Maximum temperature in the warmest month <sup>1</sup> Mean daily minimum temperature Precipitation in the driest quarter <sup>2</sup> Cloud frequency Mean wind speed

1011	Table S2. Coefficients of model-averaged general linear models of carbon stocks, gains and residence time as a function of climate, soil, continent and spatial
4040	

autocorrelation. Coefficients are AIC weighted averages across models with  $\Delta AIC < 4$  from the best performing model; variables are given a score of zero if they did not appear in a model. NA indicates that a term did not occur in any model in this set. MEM1-8 are spatial eigenvectors.

	Carbon stocks				Carbon gains				Carbon residence time			
Variable	Estimate	SE	Ζ	Р	Estimate	SE	Ζ	Р	Estimate	SE	Ζ	Р
Intercept - Africa	4.986	0.010	476.9	< 0.001	0.571	0.525	1.09	0.278	3.909	0.688	5.67	< 0.001
Minimum temperature	0.031	0.019	1.67	0.096	-0.001	0.007	0.18	0.861	0.019	0.022	0.88	0.381
Maximum temperature, warmest month	-0.089	0.022	4.11	< 0.001	-0.060	0.017	3.47	< 0.001	-0.001	0.015	0.10	0.924
Precipitation, driest quarter	0.045	0.018	2.54	0.011	-0.001	0.008	0.14	0.887	0.061	0.023	2.70	0.007
Cloud frequency	0.002	0.008	0.24	0.814	-0.006	0.011	0.54	0.592	0.025	0.021	1.17	0.241
Wind speed	0.004	0.012	0.38	0.705	0.016	0.020	0.78	0.437	-0.004	0.015	0.24	0.807
Soil texture (% clay)	0.021	0.017	1.26	0.208	-0.005	0.011	0.49	0.628	0.040	0.018	2.17	0.030
Soil fertility (CEC)	-0.003	0.009	0.34	0.732	0.005	0.011	0.51	0.613	-0.012	0.017	0.70	0.486
MEM1	0.115	0.014	7.96	< 0.001	0.319	0.559	0.57	0.569	0.375	0.734	0.51	0.610
MEM2	0.098	0.017	5.67	< 0.001	0.083	0.273	0.30	0.762	0.286	0.359	0.80	0.427
MEM3	-0.025	0.014	1.84	0.065	0.014	0.041	0.34	0.735	0.007	0.054	0.12	0.904
MEM4	-0.021	0.011	1.84	0.066	-0.038	0.020	1.84	0.066	-0.002	0.027	0.07	0.945
MEM5	0.027	0.011	2.46	0.014	0.020	0.015	1.33	0.182	0.020	0.020	0.98	0.327
MEM6	0.017	0.011	1.56	0.118	0.025	0.011	2.34	0.019	-0.014	0.014	1.05	0.293
MEM7	0.010	0.011	0.93	0.353	-0.017	0.010	1.61	0.107	0.036	0.014	2.57	0.010
MEM8	-0.072	0.013	5.64	< 0.001	0.057	0.012	4.91	< 0.001	-0.127	0.016	7.80	0.000
Asia	NA				0.380	0.542	0.70	0.485	-0.753	0.683	1.10	0.271
Australia	NA				-0.173	0.390	0.44	0.658	0.006	0.516	0.01	0.990
South America	NA				0.643	1.164	0.55	0.582	0.542	1.530	0.35	0.724

Supporting information for Sullivan et al.

1015	Table S3. Predicted biome-wide changes in long-term biomass carbon stocks (scaled to include root
1016	biomass) under global temperature increases of ~ $1.5^{\circ}$ C and ~ $2^{\circ}$ C. Changes are based on temperature
1017	effects alone, and when also accounting for the effect of increased CO <sub>2</sub> concentrations on tree growth.

- 1018 CO<sub>2</sub> effects were obtained from a synthesis of results of elevated CO<sub>2</sub> experiments (Terrer et al. (76)),
- 1019 free-air CO<sub>2</sub> enrichment (FACE) experiments (Kolby Smith et al. (73)) and CMIP5 earth system
- 1020 models (Kolby Smith et al. (73)). 95% confidence intervals around changes (based on uncertainties in
- 1021 temperature effects alone) are shown in parentheses.

CO <sub>2</sub> effect	Change in biomass carbon stocks (Pg)					
	$\sim 1.5^{\circ}$ C warming	$\sim 2^{\circ}$ C warming				
	(443 ppm CO <sub>2</sub> )	$(487 \text{ ppm CO}_2)$				
None	-26.9 (-38.415.8)	-35.3 (-49.020.9)				
Terrer et al. elevated CO <sub>2</sub> experiments	-22.0 (-33.09.9)	-26.3 (-37.611.5)				
Kolby Smith et al. FACE experiments	-6.2 (-16.8 – 7.7)	-9.9 (-24.3 – 3.9)				
Kolby Smith et al. CMIP5 models	3.9 (-8.3 – 12.6)	2.0 (-11.9 – 19.8)				